Power Calculation for the Overall Test With More Variables than Subjects

Yueh-Yun Chi^{1*}, Matthew Gribbin², and Keith E. Muller³

¹ Department of Biostatistics, University of Florida, Gainesville, FL

- ² Department of Biostatistics, Human Genome Sciences, Rockville, MD
- ³ Department of Health Outcomes and Policy, University of Florida , Gainesville, FL

* yychi@ufl.edu

This work is supported primarily by NIH/NIDCR R01-DE020832-01A1 and by NIH/NIDDK R01-DK072398, NIH/NIDCR U54-DE019261, NIH/NCRR K30-RR022258, NIH/NHLBI R01-HL091005, NIH/NIAAA R01-AA013458-01, and NIH/NIDA R01-DA031017.

WNAR, Fort Collins, CO 6/20/2012

Outline

- Motivation
- Overall test
 - Null case
 - Non-null case (power)
- Simulation
- Power analysis
- Discussion



Motivation

- Pathway analysis in microarray analysis
 - Genes are functionally or structurally related
 - System biology-driven analysis
 - Genes in the set are given a priori
 - Also important in metabolomics and proteomics
- An example
 - Wu et al. (2009) analyzed 35 sets of 4-145 genes, N=16 (9 with, 7 without metal particulate exposure)



Overall Hypothesis Testing

- One test on overall significance of the set
- # of genes (p) < sample size (N)
 MANOVA
- # of genes (p) > sample size (N)
 - Singular sample covariance matrix makes MANOVA statistics undefined
 - Use of regularization (Warton, 2008), generalized inverse (Srivastava, 2007) were proposed to "fix" MANOVA
 - A list of alternative approaches are also available



General Linear Multivariate Model (GLMM)

• Multivariate data can be modeled as follows

$$\begin{array}{ll} \boldsymbol{Y} &=& \boldsymbol{X}\boldsymbol{B} &+& \boldsymbol{E} \\ (N\times p) & (N\times q)(q\times p) & (N\times p) \end{array}$$

• Testing secondary parameters

$$H_0: \quad CBU = \Theta = \Theta_0$$
$$(a \times q)(p \times p)(p \times b) \quad (a \times b) \quad (a \times b)$$



Overall Test for Null Case

- Our alternative solution (Chi et al., 2012) is a new test for an existing statistic for GLMM
 - Controls the Type I error rate
 - Applies to general design (any GLMM)
 - Applies to data with p < N</p>
 - Easy to compute (available in SAS 9.3)



Overall Test for Null Case

• The existing statistic is a ratio of hypothesis to error sums of squares

$$t_u = [\operatorname{tr}(\boldsymbol{S}_h)/a]/[\operatorname{tr}(\boldsymbol{S}_e)/\nu_e]$$

$$oldsymbol{S}_h = (\widehat{oldsymbol{\Theta}} {-} oldsymbol{\Theta}_0)' [oldsymbol{C}(oldsymbol{X}'oldsymbol{X})^-oldsymbol{C}']^{-1} (\widehat{oldsymbol{\Theta}} {-} oldsymbol{\Theta}_0)$$

$$oldsymbol{S}_e =
u_e oldsymbol{U}' \widehat{\Sigma} oldsymbol{U}$$

 $\nu_e = N - \operatorname{rank}(\boldsymbol{X})$

Overall Test for Non-Null Case

• The exact distribution (Theorem 1, Chi et al., in preparation)

$$\Pr\{t_u \le f_0\} = \Pr\left\{\sum_{k=1}^b \pi_k y_{kh} - f_0 a \nu_e^{-1} \sum_{k=1}^b \pi_k y_{ke} < 0\right\}$$

$$y_{kh} \sim \chi^2(a, \omega_k); y_{ke} \sim \chi^2(\nu_e); y_{kh} \perp y_{ke}; \pi_k = \lambda_k / \left(\sum_{k=1}^b \lambda_k \right)$$

- A minimum set of sufficient parameters
 - Scaled variances of principal components
 - Noncentrality parameters



Overall Test for Non-Null Case

 Noncentrality parameters are 1-1 functions of the squared multiple semi-partial correlations
 Correlations have better scientific interpretability

 $\omega_k = N\rho_k^2/(1 - \rho_k^2)$ ρ_k^2 : the squared correlation between \boldsymbol{y}_{uk} and the set of predictors tested, with the predictors adjusted for all untested predictors in the model



Overall Test for Non-Null Case

• A convenient non-central *F* approximation is available (Theorem 3, Chi et al., in preparation)

$$\Pr\{t_u \le f_0\} \approx \Pr\{F(ab\epsilon_n, b\nu_e\epsilon_d, \omega_u) \le f_0\}$$

$$\epsilon_n = \left(a + 2\sum_{k=1}^b \pi_k \omega_k\right) / \left(ab\sum_{k=1}^b \pi_k^2 + 2b\sum_{k=1}^b \pi_k^2 \omega_k\right)$$

$$\epsilon_d = 1 / \left(b\sum_{k=1}^b \pi_k^2\right) = \epsilon \text{ (Sphericity parameter)}$$

$$\omega_u = \left(\sum_{k=1}^b \pi_k \omega_k\right) b\epsilon_n$$



Simulation - One Sample Problem

			Empirical, Absolute Bias in Power	
Number of Outcomes	Number of Conditions	Approximated Power	Max	Mean
$ \begin{array}{r} 64 \\ 64 \\ 64 \\ 64 \end{array} $	$36 \\ 36 \\ 36 \\ 36 \\ 36$	$egin{array}{c} 0.20 \\ 0.50 \\ 0.80 \\ 0.90 \end{array}$	$\begin{array}{c} 0.010 \\ 0.024 \\ 0.005 \\ 0.013 \end{array}$	$\begin{array}{c} 0.004 \\ 0.004 \\ 0.002 \\ 0.002 \end{array}$
$\begin{array}{c} 256 \\ 256 \end{array}$	$\begin{array}{c} 36\\ 36\end{array}$	$\begin{array}{c} 0.80\\ 0.90\end{array}$	$\begin{array}{c} 0.005 \\ 0.005 \end{array}$	$\begin{array}{c} 0.002\\ 0.002\end{array}$
$\frac{1024}{1024}$	$\frac{36}{36}$	0.80 0.90	0.004 0.003	$0.001 \\ 0.001$

 $N \in \{10, 20, 40\}, \epsilon \in \{0.27, 0.56, 0.76\}$, number of nonzero ρ_k^2 of either 4 or 32, Location of nonzero ρ_k^2 at either the most dominant or middle components



Simulation - Two Sample Problem

			Empirical, Absolute Bias in Power	
Number of Outcomes	Number of Conditions	Approximated Power	Max	Mean
64	36	0.20	0.001	0.006
64	36	0.50	0.034	0.011
64	36	0.80	0.037	0.011
64	36	0.90	0.028	0.009
256	36	0.80	0.031	0.010
256	36	0.90	0.024	0.009
1024	36	0.80	0.027	0.011
1024	36	0.90	0.022	0.009

 $N \in \{10, 20, 40\}, \epsilon \in \{0.27, 0.56, 0.76\}$, number of nonzero ρ_k^2 of either 4 or 32, Location of nonzero ρ_k^2 at either the most dominant or middle components



Power Analysis

- Seven input components
 - Type I error rate
 - Design matrix
 - Between-subject contrast matrix
 - Within-subject contrast matrix
 - Null matrix
 - Primary parameters matrix
 - Error covariance matrix



Power Analysis

• Number of parameters explodes when p >> N

Type I error rate	lpha	1×1
Design matrix	X	$N \times q$
Between-subject contrast matrix	\boldsymbol{C}	$a \times q$
Within-subject contrast matrix	${oldsymbol U}$	p imes b
Null matrix	$\mathbf{\Theta}_0$	$a \times b$
Primary parameters matrix	\boldsymbol{B}	q imes p
Error covariance matrix	$\boldsymbol{\Sigma}$	p imes p
		-



14

Power Analysis

• Power equivalence simplifies problem (Thm. 2)

	Hypothesis Testing Scenario				
Feature	S_1	S_2			
Model	$Y_1 = X_1 B_1 + E_1$	$m{Y}_2 = m{X}_2 m{B}_2 + m{E}_2$			
# of Outcomes	p	b			
# of Predictors	q	$r = \operatorname{Rank}(\boldsymbol{X}_1)$			
Design Matrix Feature	$oldsymbol{X}_1$	$oldsymbol{X}_2'oldsymbol{X}_2=oldsymbol{I}_r$			
Error Covariance	$\mathcal{V}(oldsymbol{E}_1) = oldsymbol{\Sigma}$	$\mathcal{V}(oldsymbol{E}_2) = \mathrm{Dg}(oldsymbol{\pi})$			
Between-Subject Contrast	t C_1	$oldsymbol{C}_2 = \left[oldsymbol{I}_a oldsymbol{0} ight]$			
Within-Subject Contrast	$oldsymbol{U}_1$	$oldsymbol{U}_2 = oldsymbol{I}_b$			
Primary Parameters	$oldsymbol{B}_1$	$oldsymbol{B}_2 = \left[oldsymbol{eta}_2 oldsymbol{0} ight]'$			
$eta_{2k} = (\pi_k \omega_k)^{1/2}, \pi_k = \lambda_k / \left(\sum_{k=1}^b \lambda_k \right)$					

F

Conclusion for *p* > *N*

- Overall testing is important for pathway analysis
- We have
 - $\ ^{\rm o}$ A size α general test for null case
 - Accurate power approximation
 - Software for testing available
 - Updated power software underway
- Power equivalent scenarios help simplify the power analysis



16

Discussion

- With high outcome dimension, practicing safe computing is particularly essential to ensure numerical accuracy
- Power calculation with random covariate will be explored for future research



Reference

- Chi et al. (2012) Statistics in Medicine, in press
- Srivastava (2007) Journal of Japanese Statistical Society, 37, 53-86
- Warton (2008) JASA, 103, 340-349
- Wu et al. (2009) Bioinformatics, 25, 1145-1151



18